

MATH 4281 Risk Theory–Ruin and Credibility

Start Module 2: Ruin Theory

Feb 2, 2021

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Recap and Motivation

The Story so Far

- As it stands now- the models we have studied have assumed a **short time frame**.
- We have tried to model aggregate claims over say a week/month/year.
- Assumed no **Time Value of Money** or rigorous model for **Premiums**.

Some Questions

Q1 What happens if we can't pay all the claims?

Q2 How do we set premiums to guarantee that we can?

Q3 How does **Time** factor in to this?

These are the questions that we will explore in this module

Stochastic Processes

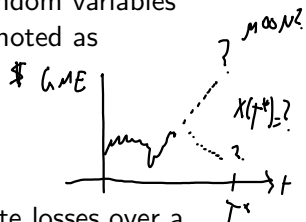
What is a Stochastic Process

- Stochastic - from the Greek for "to aim" or "to guess". Generally adjective denoting "randomness" e.g:
 - stochastic process (mathematics)
 - stochastic resonance (biology)
 - newsworthy "stochastic terrorism" (social sciences)
 - etc...
- Process - Latin for "progression"
- Stochastic Process - stands to reason this is some progression of random events

What is a Stochastic Process

- A *stochastic process* is any collection of random variables $X(t)$, $t \in T$. This stochastic process is denoted as

$$\{X(t), t \in T\}.$$



- We are interested in modelling the aggregate losses over a given period of time, not necessarily **at one point**!
- For example: the aggregate loss process denoted by $\{S(t), t \geq 0\}$, where $S(t)$ is the aggregate loss **at time t** .

Independent Increments

A stochastic process $\{X(t), t \geq 0\}$ has *independent increments* if:

- For all $t_0 < t_1 < t_2 < \dots < t_n$ the following RVs¹ are independent:

$$\underbrace{X(t_1) - X(t_0)}_{Y_1}, \underbrace{X(t_2) - X(t_1)}_{Y_2}, \dots, \underbrace{X(t_n) - X(t_{n-1})}_{Y_n}$$

- That is, future increases are independent of the past.

¹RV = Random Variable

Stationary Increments

A stochastic process $\{X(t), t \geq 0\}$ has *stationary increments* if:

- for all choices of t_1, t_2 and $\tau > 0$:

$$X(t_2 + \tau) - X(t_1 + \tau) \stackrel{d}{=} X(t_2) - X(t_1)$$

- Equivalently for $s < t$

$$X(t) - X(s) \stackrel{d}{=} X(t - s)$$

Counting process

- A stochastic process $\{N(t), t \geq 0\}$ is a *counting process* if it represents the number of events that occur up to time t .
- Q: What is the significance of counting processes?
- A: We will use them to model the number of claims received during a particular time.

Counting process

A counting process $\{N(t), t \geq 0\}$ must satisfy:

- 1 $N(t) \geq 0$.
- 2 $N(t)$ is integer-valued.
- 3 $N(s) \leq N(t)$ for any $s < t$, i.e. it must be non-decreasing.
- 4 For $s < t$, $N(t) - N(s)$ is the number of events that have occurred in the interval $(s, t]$.

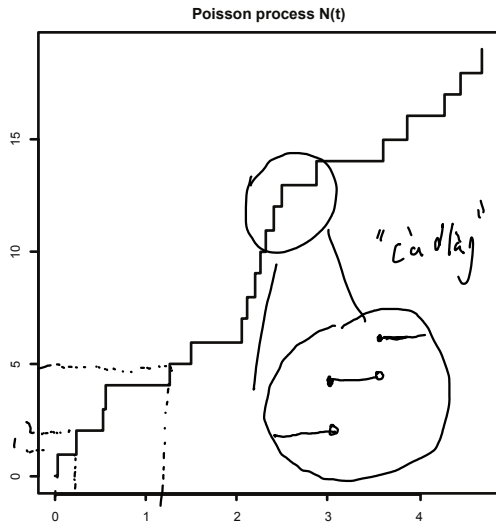
Review

A counting process $\{N(t), t \geq 0\}$ is a *Poisson process* with rate λ , for $\lambda > 0$, if:

- 1 $N(0) = 0$;
- 2 it has independent increments; and
- 3 the number of events in any interval of length t has a Poisson distribution with mean λt . That is, for all $s, t \geq 0, n = 0, 1, \dots$

$$\Pr[N(t+s) - N(s) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}.$$

A path (realization) of the Poisson process



- Counting process
- Step function
- What is the arrival time?

A Noteworthy Characterization

Theorem:

- Consider the time from the $(i-1)$ th and i th jump W_i .
- That is $t = W_1 + \dots + W_{N(t)}$
- Then $N(t+h) - N(t) \sim \text{Poi}(\lambda h)$ iff $W_i \sim \text{Exp}(1/\lambda)$

Proof:

$$\begin{aligned} (W_i \leq \tau) &\equiv (N(t_i + \tau) - N(t_i) > 0) \\ P(W_i \leq \tau) &= P(N(t_i + \tau) - N(t_i) > 0) \quad \text{complement} \\ P(W_i \leq \tau) &= 1 - P(N(t_i + \tau) - N(t_i) = 0) \\ &= 1 - \frac{(\lambda \tau)^0}{0!} e^{-\lambda \tau} = 1 - \underbrace{e^{-\lambda \tau}}_{\text{Exp}(1/\lambda)} \end{aligned}$$

A Noteworthy Characterization

- Recall also that there is something special about the distribution!
- Exponential waiting times are **Memoryless**

E.g:

$$\hookrightarrow P(W > t+s \mid W > t) = P(W > s)$$

$$\frac{P(W > t+s)}{P(W > t)} = P(W > s) \Rightarrow \underbrace{S(t+s)}_{\substack{\text{Exponential Survival Function is only} \\ \text{soln to the functional equation above}}} = \underbrace{S(t)}_{\substack{\text{Exponential Survival Function is only} \\ \text{soln to the functional equation above}}} \underbrace{S(s)}_{\substack{\text{Exponential Survival Function is only} \\ \text{soln to the functional equation above}}}$$

$s \rightarrow$ survival fn

Zooming in on the process

Explain why the following are true:

$$\begin{aligned}
 P[N(t+dt) - N(t) = 1 | N(s), 0 \leq s \leq t] &= \lambda dt + o(dt) \\
 P[N(t+dt) - N(t) = 0 | N(s), 0 \leq s \leq t] &= 1 - \lambda dt + o(dt) \\
 P[N(t+dt) - N(t) \geq 2 | N(s), 0 \leq s \leq t] &= o(dt)
 \end{aligned}$$

$$i) P(N(t+dt) - N(t) = 1) = \frac{e^{-(\lambda dt)} (\lambda dt)^1}{1!}$$

$$\begin{aligned}
 &= \lambda dt [1 - (-\lambda dt) + o(dt)] \\
 &= \lambda dt + \underbrace{\lambda^2 dt^2 + \dots}_{o(dt)}
 \end{aligned}$$

N.B

$$f(x) \in O(g(x)) \text{ if}$$

$$\lim_{x \rightarrow 0} \frac{f(x)}{g(x)} = 1$$

ii)

$$P(N(t+dt) - N(t) = 0) = \cancel{(\lambda dt)^0} e^{-(\lambda dt)}$$

$\swarrow$
 $\searrow$
 1

$$= [1 - \lambda dt + o(dt^2)]$$

$$\begin{aligned} \text{iii)} \quad & P(N(t+dt) - N(t) \geq 2) \\ &= 1 - P(N(t+dt) - N(t) < 2) \\ &= \dots ? \quad (\text{Exercise}) \end{aligned}$$

Properties of the Poisson process: Summary

If $\{N(t), t \geq 0\}$ is a *Poisson process* with rate λ , for $\lambda > 0$, then

- 1 $N(0) = 0$;
- 2 it has independent **and** stationary increments;
- 3 It can never have more than 1 jump at a time! That is:

$$\Pr[N(t+h) - N(t) = 0] = e^{-\lambda h} = 1 - \lambda h + o(h)$$

$$\Pr[N(t+h) - N(t) = 1] = \lambda h e^{-\lambda h} = \lambda h + o(h)$$

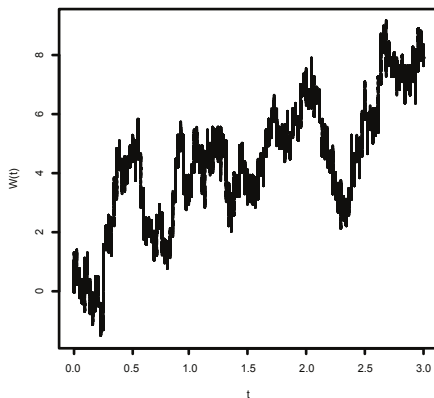
and

$$\Pr[N(t+h) - N(t) \geq 2] = o(h)$$

- 4 The time between two consecutive jumps follows the $\text{Exponential}(\lambda)$ distribution.

Brownian motion as the limit of a shifted Poisson process

Approximation of a Brownian Motion



($\mu = 2$, $\sigma = 5$, and $\tau = 0.02$)

Consider the following shifted Poisson process:

$$W(t) = \tau N(t) - ct.$$

Increments have moments

$$E[W(t+h) - W(t)] = (\tau\lambda - c)h \equiv \mu h,$$

$$\text{Var}(W(t+h) - W(t)) = (\tau^2\lambda)h \equiv \sigma^2 h$$

When $\tau \rightarrow 0$ for fixed μ and σ^2 ,
 $\{W(t)\}$ becomes a Brownian motion with parameters μ and σ^2 .

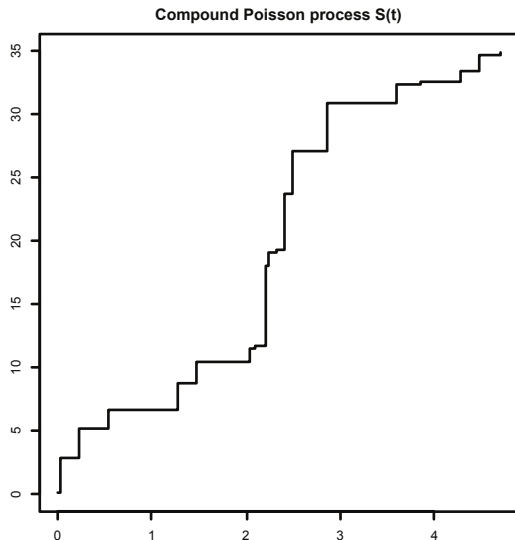
We define a **Compound Poisson process** $\{S(t), t \geq 0\}$ like so:

$$S(t) = \sum_{i=1}^{N(t)} X_i.$$

Where:

- $\{N(t)\}$ is a Poisson process with parameter λ
- $\{X_i\}$ are iid $\sim P(x)$

A path (realization) of the compound Poisson process



- Now step i has height X_i instead of 1.
- Increments
 $S(t+h) - S(t) \sim$
Compound
Poisson($\lambda h, P(x)$)

Mean and Variance of the compound Poisson process:

$$E[S(t)] = \lambda t E[X], \quad \text{Var}[S(t)] = \lambda t E[X^2].$$

Law of total prob/Tower Property

$$\begin{aligned} & E[E[S(t) \mid N(t) = n]] \\ &= E[E[\sum_{i=1}^N X_i \mid N(t) = n]] \\ &= E[N(t)] E[X_i] \\ &= (\lambda t) E[X_i] \end{aligned}$$

Same but law of total Var
(Exercise)

The MGF of the compound Poisson process:

$$M_{S(t)}(z) = \exp\{\lambda t[M_X(z) - 1]\}.$$

Pf:

$$\begin{aligned} & E\left[e^{z S(t)} \right] \\ &= E\left[E\left[e^{z \sum_{i=1}^{N(t)} X_i} \mid N(t) \right] \right] \\ &= E\left[\left(M_{X_i}(z) \right)^{N(t)} \right] \\ &= E\left[e^{N(t) \ln(M_{X_i}(z))} \right] = M_{N(t)}(\ln(M_{X_i}(z))) \end{aligned}$$