

MATH 4281 Risk Theory—Ruin and Credibility

Module 2: Ruin Theory (cont.)

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Post Reading Week/test #1 Review

Recall the following

Since it's been since the 4th we will review what we covered of Ruin Theory so far. Recall we covered:

- Stochastic processes and their properties (independent \ stationary increments, etc...) .
- Counting processes, specifically Poisson processes.
- Compound Poisson processes.
- Leading to the **Cramér-Lundberg process**:

$$U(t) = \underbrace{u_0 + ct}_{\text{Revenue}} + \underbrace{\sum_{i=1}^{N(t)} X_i}_{\text{Losses}}$$

i.e.

The probability of ruin

- Recall the Cramér-Lundberg model:

$$U(t) = u_0 + ct - \sum_{i=1}^{N(t)} X_i$$

- The time to ruin T is defined as

$$T = \inf\{t \geq 0 \mid U(t) < 0\}.$$

- The probability that the company would be ruined by time t is denoted by

$$\psi(u_0, t) = \Pr[T < t].$$

Avoiding Ultimate Ruin

- Finally, the probability of **ultimate** ruin is

$$\psi(u_0) = \Pr(T < \infty) = \lim_{t \rightarrow \infty} \psi(u_0, t) \geq \psi(u, t).$$

- The Net Profit Condition (NPC):

$$c \leq \lambda \mathbb{E}[X_i] \Rightarrow \psi(u_0) = 1$$

- To ensure the NPC holds we add our "safety loading" :

$$c = \underline{(1 + \theta)\lambda \mathbb{E}[X]}$$

Recall we introduced an approximation

We can approximate ψ easy via **The Lundberg Inequality**:

$$\psi(u) \leq e^{-Ru}$$

Where R (the adjustment coefficient) solves the equation¹

$$e^{rpt} = \mathbb{E}[e^{rS_t}]$$

Today we will discuss this in more detail.

¹Where $S_t = \sum_{i=1}^{N(t)} X_i$

The Lundberg Inequality

Avoiding Ultimate Ruin

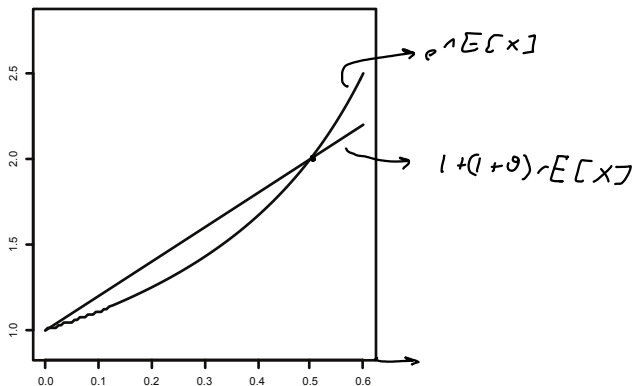
In the Cramér-Lundberg model, consider the excess of losses over premiums over the interval $[0, t]$: $S(t) - ct$. We define the **adjustment coefficient R** as the first positive solution of the following equation in r :

$$M_{S(t)-ct}(r) = E \left[e^{r(S(t)-ct)} \right] = e^{-rct} e^{\lambda t [M_X(r)-1]} = 1,$$

Recall $c = (1 + \theta)\lambda E[X]$. So, the adjustment coefficient R is the first positive of the following equation:

$$1 + (1 + \theta)rE[X] = M_X(r)$$

Does such an R exist?



- Recall that Jensen's inequality gives $1 + (1 + \theta)rE[X] = M_X(r) \geq e^{rE[X]}$
- How else could this fail?

The Theorem

- ① Let $R > 0$ be the adjustment coefficient. If $\{U(t)\}$ is a Cramér-Lundberg process with $\theta > 0$, then for $u \geq 0$

$$\psi(u) = \frac{e^{-Ru}}{E[e^{-RU(T)} | T < \infty]}.$$

- ② Since $U(T) < 0$, we have then (Lundberg's exponential upper bound)

$$\psi(u) < e^{-Ru}.$$

An Example

Assume $X \sim \exp(\beta)$ (the mean is $1/\beta$). Find R and $\psi(u)$.

1) R

$$e^{rc} = E[e^{rS_1}]$$

$$e^{rc} = \exp\{\lambda + (M_X(r) - 1)\}$$

$$\Rightarrow rc = \lambda (M_X(r) - 1)$$

$$1 + r(c/\lambda) = \frac{1}{1 - r/\beta}$$

$$[1 + r(c/\lambda)] \cdot [1 - r/\beta] = 1$$

$$1 + (c/\lambda - 1/\beta)r - (c/\lambda)(1/\beta)r^2 = 1$$

$$\Rightarrow r = \beta - \lambda c > 0$$

Recall

$$M_X(r) = \frac{\beta}{\beta - r} \quad (MGR)$$

NPC

$$c > \lambda E[X]$$

$$c > \lambda/\beta$$

$$\lambda c < \beta$$

An Example

$$\begin{aligned}\text{So then } R &= B - \frac{\lambda}{c} \\ &= B - \frac{\lambda}{(1+\theta)\lambda(1/B)} \\ &= B \frac{\theta}{1+\theta}\end{aligned}$$

$$\Rightarrow \boxed{\psi(u) \leq \exp\left\{-B \frac{\theta}{1+\theta} u\right\}}$$

$$(N.B. \psi(u) = \frac{1}{1+\theta} \exp\left\{-B \frac{\theta}{1+\theta} u\right\})$$

Almost never
got a "nice"
formula like
this!

Proving the Cramér-Lundberg Inequality

Proving our theorems

To start show that $\{e^{-RU(t)}\}$ is a martingale²

→ adj. cor.

→ X_t is a s.p. then

$$E[X_t | X_{t-1}] = X_{t-1}$$

$$E[e^{-RU(t)} | e^{-RU(s)}] \quad s < t$$

$$= E\left[\exp\left\{-R\left(u_0 + c(\overbrace{t-s}^{+s}) + \sum_{i=1}^{N(s)} X_i + \sum_{i=N(s)+1}^{N(t)} X_i\right)\right\} \mid e^{-RU(s)}\right]$$

$$= E\left[\underbrace{\exp\left\{-R\left(u_0 + cs + \sum_{i=1}^{N(s)} X_i\right)\right\}}_{= e^{-RU(s)}} \exp\left\{-R\left(c(t-s) + \sum_{i=N(s)+1}^{N(t)} X_i\right)\right\} \mid e^{-RU(s)}\right]$$

$$= c(t-s) + \sum_{i=1}^{N(t-s)} X_i$$

$$\Rightarrow E\left[\exp\left\{-R(c(t-s) - S_{t-s})\right\}\right] = 1$$

Hence the definition of "R"!

$$\rightarrow \text{i.e. } E[e^{-RU(t)} | e^{-RU(s)}] = e^{-RU(s)} \text{ for } s < t \text{ or } E[e^{-RU(t)}] = e^{-Ru} \text{ for all } t$$

A very very useful theorem

(Given we don't have all the machinery we need at this point- we will define a *stopping time* as a random time dependent on another stochastic process exhibiting some behaviour)

Theorem (Optimal Stopping Theorem)

Given a bounded stopping time T , i.e. $T \leq t_0 < \infty$ for a martingale^a M_t them:

$$M_0 = E[M_T]$$

^aFor those who know we must also impose right continuity

An example: Gambler's Ruin

A gambler enters a casino with n dollars and plays a game with a win probability p . He gains \$1 for every win and losses \$1 for every loss. He leaves when he wins N or loses everything. What is the probability he leaves ruined?

X_t - gambler's wealth

$$f(X_t) = \lambda^{X_t} \quad \text{where} \quad \lambda = \frac{1-p}{p}$$

$$\Rightarrow E[f(X_{t+1}) | \mathcal{F}(X_t)] = p \lambda^{(X_t+1)} + (1-p) \lambda^{(X_t-1)} = \lambda^{X_t} = f(X_t)$$

Hence $f(X_t)$ is a Martingale!

An example: Gambler's Ruin

$$\tau = \inf_{t \geq 0} \{ X_t = 0 \text{ or } N \} \rightarrow \text{stopping time}$$

by O.S.T

$$\underbrace{E[f(X_0)]}_{f(\lambda)} = E[f(X_\tau)] = P(X_\tau = N) f(N) + \underbrace{P(X_\tau = 0)}_{1 - P(X_\tau = N)} f(0)$$

$$\Rightarrow \boxed{P(X_\tau = N) = \frac{1 - f(\lambda)}{1 - f(N)} = \frac{1 - \lambda^N}{1 - \lambda^N}}$$

Proof of the Cramér-Lundberg Inequality

- So we can do something similar with $e^{-Ru(x)}$
Now we are interested in $T = \inf_{t \geq 0} \{U_t < 0\}$

$$\underline{P, U, \text{ or } P(T = \infty) > 0 \text{ i.e. unbounded}}$$

take $t_0 \leq \infty$; $Z = T \wedge t_0 = \min\{T, t_0\}$ is bounded

Apply O.S.T;

$$\begin{aligned} e^{-R(U(0))} &= E[e^{-RU(Z)}] \\ e^{-Ru} &= E[e^{-RU(Z)} | T < t_0] P(T < t_0) \\ &\quad + E[e^{-RU(Z)} | T \geq t_0] P(T \geq t_0) \end{aligned}$$

Proof of the Cramér-Lundberg Inequality

$$e^{-Ru} = E[e^{-Ru_\tau} \mid T < t_0] \overbrace{P(T < t_0)}^{\psi(u)} + E[e^{-Ru_\tau} \mid T \geq t_0] P(T \geq t_0)$$

$$\begin{aligned} E[e^{-Ru_\tau} \mid T \geq t_0] &= E[e^{-Ru(\tau)} \mathbb{1}_{\{T \geq t_0\}}] \\ &= E[e^{-Ru(\tau)} \mathbb{1}_{\{u(\tau) \geq 0\}}] \end{aligned}$$

Since $0 \leq e^{-Ru(\tau)} \mathbb{1}_{\{u(\tau) \geq 0\}} \leq 1$

$$\Rightarrow e^{-Ru(t)} \mathbb{1}_{\{u(t) \geq 0\}} \longrightarrow 0 \text{ as } t_0 \longrightarrow \infty$$